

An Analysis of Change of Major Behavior of Cal Poly Students

A Senior Project

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Introduction

There has been much work done in educational research that attempts to model and explain various student behaviors from a multitude of perspectives. Much of the research done on university students focuses on trying to understand factors associated with college degree attainment. This is of no surprise as it is a fairly common belief that a college education is important for the individual and is inextricably linked to the health of the national economy. Further, graduation rates are a common measurement of the success of a university.

Educational researchers in the past have investigated various background and personality characteristics as well as institutional factors that are significantly associated with college graduation rates, but little work has been done to assess the impact that changing majors has on degree attainment. In particular, there has been very little work estimating the relationship between switching majors and time until degree attainment, especially using survival analysis methods. The main goal of this study was to present a reasonable way with which to accomplish such an analysis. Though the end goal is to be able to answer questions about the relationship between switching major and college degree attainment on a national level, performing the relevant analyses on the data available from Cal Poly San Luis Obispo was used to assess the effectiveness of such an approach and provide a good starting point for future research.

Before looking into the Cal Poly data, it is advantageous to first take a look at what results other researchers have found regarding factors related to graduation from college as well as any analyses of major switching behaviors.

Literature Review

Chen and Weko (2009) focus their attention on developing a profile of undergraduate students who enter and subsequently graduate from STEM majors. They analyzed longitudinal national level data and conducted simple t-tests without adjusting for multiple comparisons but they did obtain some interesting results. They did not detect any gender differences in STEM degree attainment, though they noted that more men entered STEM majors. They found that Asian/Pacific Islander students were much more likely to enter a STEM major than White, African American, and Hispanic/Latino students, between which no measurable differences were discovered. However, both White and Asian/Pacific Islander students were more likely than African American or Hispanic/Latino students to graduate from a STEM major. Younger students and those considered dependent (usually under age 24) had a higher STEM major entrance rates and STEM graduation rates. Students having at least one parent with a four year degree and students with considerable academic preparation also were shown to have higher STEM graduation rates. Overall, students entering college in STEM majors were determined to be more likely to graduate from college. An interesting aspect of their analysis was that they broke all STEM majors into four groups: Mathematics, the natural sciences, engineering, and computer science. They conducted analyses that sought to determine if there were different graduating behaviors for the various STEM types and did indeed find several important differences.

Shaw and Barbuti (2010) looked at factors that are associated with a student remaining in their original major through the third year of college with an explicit focus on STEM majors. Using national level data collected from 39 four year institutions, they found several interesting relationships between various background and demographic factors and major persistence through the third year as well as academic performance measures while in college. Statistical significance was assessed using Cohen's d, a method that standardizes mean differences between groups and uses this as a measure of effect size. Though their study only looks at persistence in matriculating major through the third year, they do find some interesting results. They found small differences in matriculating major persistence through the third year in gender, parental income, and first generation status, as well as with ethnicity. They found that women were less likely to persist than men, underrepresented ethnicities were less likely to persist than White and Asian/Pacific Islander students, and first generation students were less likely than students whose parents had some college to persist. They also note that students that persisted in their matriculating major through the third year also tended to have higher high school and university grade point averages (GPA) but this varied by majors.

Allen and Robbins (2007) attempted to predict persistence to the third year in matriculating major using academic preparation measures, first year academic performance, and the student's vocational interests. Using logistic regression methods they determined that first year GPA was associated with persistence but that neither high school GPA nor ACT composite score was significant in the presence of first year GPA. They also note that their finding of the importance for first year college GPA is consistent with results obtained by other researchers.

Chizmar (2010) published his analysis of the role that gender plays in the persistence of economics majors in a paper published by *The Journal of Economic Education* in 2000. Though focused on specifically economics he failed to detect a gender difference but was able to conclude that students

whose grades in economics classes were lower than their grades in other classes were more likely to switch as well as students with only a small amount of coursework completed in the major. More interesting than his results was his method of analysis. He made use of the discrete time hazard model as formulated by researcher Judith Singer and John Willett of Harvard. This is notable as it is a method that we use to analyze our data and will be discussed at length later.

Methods

Work on the project began in Summer 2011. The first big task undertaken was to clean the data, create and/or modify variables, and recode them with meaningful labels. All of the statistical analysis and manipulation for the project was done using the R statistical software package.

Research Questions

Four main questions were developed to study the relationship between switching majors, graduating from college, and time until graduation.

- (1) Can we identify important demographic and academic characteristics associated with the chance that a student will change his/her major?
- (2) Can we determine, after controlling for background information and academic performance (both during high school and while at Cal Poly), what effect does switching major have on the risk of graduation?
- (3) Can we determine if students that switch majors have different graduation rates than those that do not change majors?
- (4) Can we determine how the timing of the major switch affects time until graduation?

The Data, Variables, and Sample Used

The data were acquired from Institutional Planning and Analysis at Cal Poly San Luis Obispo and provides information about the incoming freshman class of 2005. This data set includes academic information about their studies at Cal Poly for 6 years beginning Fall Quarter 2005 through Spring Quarter 2011, as well as various background information concerning ethnicity, gender, parental education, California residency and high school academic performance. A comprehensive list of the variables used, their levels, and explanations are located in Table 1.

The type of information provided to us by the variables fell into three classes. The first class of variables capture demographic background information. These include *Gender*, *Ethnicity*, *First Generation Status*, *Geography*, and *Pell Grant*. The second class of variables contain academic background information that perhaps partly capture the preparedness of the student for the rigors of Cal Poly coursework. These include *High School GPA*, *SAT Score*, and *Remedial Work*. The final class of variables include information about each student's Cal Poly career: the units and term GPA they had for every quarter from Fall 2005 to Summer 2011, what major/college they belonged to, and major switching information.

Table 1: Variables Used in Analysis

| Variable | Levels | Additional Information |
|---|--|---|
| <i>Demographic Background Characteristics:</i> | | |
| <i>Gender</i> | Male Female | |
| <i>Ethnicity</i> | Hispanic/Latino African American Native American White Haw/ Pacific Islander Asian American International Unknown | International students and students of unknown ethnicity were not used in inferential models. |
| <i>FirstGeneration</i> | Student is First Gen Student is not First Gen | A Student is considered to be a First Generation Student if the maximum educational level of either parent was high school or below. |
| <i>Geography</i> | CA Resident non-CA Resident | A student is a CA Resident if they were a resident at the time they applied to Cal Poly. |
| <i>Pell Grant</i> | Received Pell Grant No Pell Grant | A Pell Grant is a form of Financial Aid provided by the government. |
| <i>Academic Background Characteristics:</i> | | |
| <i>High School GPA</i> | | The final high school GPA, on a 4.0 scale. |
| <i>SAT Score</i> | | The original data contained information about each student's SAT and ACT scores. Not all students took both entrance exams. For those students that only took the ACT, an SAT equivalence score was computed. The SAT score used was the sum of the SAT reading and SAT mathematics scores. computed using a concordance table. |

| <u>Variable</u> | <u>Levels</u> | <u>Additional Information</u> |
|---|---|--|
| <i>Remed</i> | Did some remedial work No remedial work required | When entering the university students take mathematics and english placement exams to see if they are ready for college level work. |
| <i>Academic Performance at Cal Poly:</i> | | |
| <i>Units</i> | | Number of units attempted for each quarter from 2005-2011. |
| <i>GPA</i> | | Term GPA for each quarter from 2005-2011. |
| <i>College</i> | CAFES CAED CENG CLA OCOB CSM | The college the student's major belonged to upon entering Cal Poly in 2005. <i>College of Architecture and Engineering Science</i> <i>College of Education</i> <i>College of Engineering</i> <i>College of Liberal Arts</i> <i>Orfalea College of Business</i> <i>College of Science and Mathematics</i> |
| <i>STEM Matric</i> | STEM non-STEM | All majors at Cal Poly were classified as being a STEM or a non-STEM major based on course work required. This variable refers to the classification of the student's matriculating major. |
| <i>FirstSwitch.Vary</i> | | This is a variable we created to keep track of type and timing of a student's first major switch. This variable is coded as "Persist" for all quarters before the student's first major change, is either a 1 or a 2 for all quarters after the first major switch. A "1" indicates that the major change was to a STEM major, "2" indicates the change was to a non-STEM major. |
| <i>First Switch</i> | Persist Switch to STEM Switch to non-STEM | This variable tracks if a major change occurred. The coding only applies to the first major change. |
| <i>Number of Major Changes</i> | | This variable represents the number of major changes made by the student from 2005-2011. |

| <u>Variable</u> | <u>Levels</u> | <u>Additional Information</u> |
|------------------------|--|---|
| <i>Year Switch</i> | Persist Switch Year 1 Switch Year 2 Switch Year 3 Switch Year 4 Switch Year 5 | This variable indicates the academic year in which the first major switch was made. |

Descriptive Statistics

We are specifically interested in looking at the association between switching majors and overall college graduation rates. To begin, we take a look at some descriptive statistics that give us a feel for what the overall switching and graduation behaviors are.

In Table 2, we see that of the original 3425 students, 2596 managed to graduate within six years and 829 had yet to complete their degrees. We see that of those that managed to graduate in the six years, about 24.4% switched majors compared to 12.1% for those that have not yet graduated. Looking more specifically at the type of major switch, we see that of the 21.4% of students that switched majors at some point, 9.9% switched to a STEM major and 11.5% switched to a non-STEM major. In addition, we see that the relative percentage contribution of those that switched to STEM majors and those that non-STEM majors are higher to the Graduation group and lower for the Drop Out/Still Enrolled group. This is encouraging as it seems to point to higher graduation rates for those that switch majors. Also note, we see that the overwhelming majority of students that did change majors did so only once, and that it is fairly rare that a student change majors two or more times, only about 1.3% of all students did so.

Table 2: Major Change Information Broken Down by Graduation Status for 2005 Cal Poly Freshman Class

| Variable | Levels | nDropOut/StillEnrolled | %DropOut/StillEnrolled | nGrad | %Grad | nall | %all |
|---------------------------|--------------------|-------------------------------|-------------------------------|--------------|--------------|-------------|-------------|
| EverChange | Persist | 729 | 87.9 | 1964 | 75.7 | 2693 | 78.6 |
| | Switched | 100 | 12.1 | 632 | 24.4 | 732 | 21.4 |
| | all | 829 | 100.0 | 2596 | 100.0 | 3425 | 100.0 |
| First Major Switch | Persist | 729 | 87.9 | 1964 | 75.7 | 2693 | 78.6 |
| | Sw STEM | 61 | 7.4 | 278 | 10.7 | 339 | 9.9 |
| | Sw Non-STEM | 39 | 4.7 | 354 | 13.6 | 393 | 11.5 |
| | all | 829 | 100.0 | 2596 | 100.0 | 3425 | 100.0 |
| NumberMajorChanges | 0 | 729 | 87.9 | 1964 | 75.7 | 2693 | 78.6 |
| | 1 | 90 | 10.9 | 597 | 23.0 | 687 | 20.1 |
| | 2 | 8 | 1.0 | 34 | 1.3 | 42 | 1.2 |
| | 3 | 1 | 0.1 | 1 | 0.0 | 2 | 0.1 |
| | 4 | 1 | 0.1 | 0 | 0.0 | 1 | 0.0 |
| | all | 829 | 100.0 | 2596 | 100.0 | 3425 | 100.0 |

Now, let's look at the graduation and major switching rates tabulated by specific categorical variables, this information is available in Table 3.

Table 3: Background Characteristics Broken Down by Graduation Status for 2005 Cal Poly Freshman Class

| Variable | Levels | nGraduated | %Graduated | nDropOut/StillEnrolled | %DropOut/StillEnrolled | nall | %all |
|-----------------|-------------------|------------|------------|------------------------|------------------------|------|-------|
| College | CAED | 234 | 9.0 | 72 | 8.7 | 306 | 8.9 |
| | CAFES | 483 | 18.6 | 171 | 20.6 | 654 | 19.1 |
| | CENG | 520 | 20.0 | 268 | 32.3 | 788 | 23.0 |
| | CLA | 449 | 17.3 | 108 | 13.0 | 557 | 16.3 |
| | CSM | 294 | 11.3 | 100 | 12.1 | 394 | 11.5 |
| | OCOB | 616 | 23.7 | 110 | 13.3 | 726 | 21.2 |
| | all | 2596 | 100.0 | 829 | 100.0 | 3425 | 100.0 |
| Ethnicity | AfAm | 22 | 0.8 | 11 | 1.3 | 33 | 1.0 |
| | AsAm | 292 | 11.2 | 87 | 10.5 | 379 | 11.1 |
| | Haw | 8 | 0.3 | 7 | 0.8 | 15 | 0.4 |
| | Hisp | 214 | 8.2 | 105 | 12.7 | 319 | 9.3 |
| | Inter | 12 | 0.5 | 5 | 0.6 | 17 | 0.5 |
| | NaAm | 24 | 0.9 | 7 | 0.8 | 31 | 0.9 |
| | Unknown | 228 | 8.8 | 85 | 10.2 | 313 | 9.1 |
| | White | 1796 | 69.2 | 522 | 63.0 | 2318 | 67.7 |
| | all | 2596 | 100.0 | 829 | 100.0 | 3425 | 100.0 |
| Gender | Male | 1261 | 48.6 | 524 | 63.2 | 1785 | 52.1 |
| | Female | 1335 | 51.4 | 305 | 36.8 | 1640 | 47.9 |
| | all | 2596 | 100.0 | 829 | 100.0 | 3425 | 100.0 |
| FirstGeneration | FirstGen | 2397 | 92.3 | 721 | 87.0 | 3118 | 91.0 |
| | NotFirstGen | 199 | 7.7 | 108 | 13.0 | 307 | 9.0 |
| | all | 2596 | 100.0 | 829 | 100.0 | 3425 | 100.0 |
| Remed | AtLeastSomeREMEDI | 282 | 10.9 | 156 | 18.8 | 438 | 12.8 |
| | NoREMEDI | 2314 | 89.1 | 673 | 81.2 | 2987 | 87.2 |
| | all | 2596 | 100.0 | 829 | 100.0 | 3425 | 100.0 |

For the *College* variable, we note that most colleges have roughly equal contributions (in terms of percentage) to both the DropOut/Still Enrolled and Graduation groups. However, we see that 32.3% of all Cal Poly students that did not graduate in the six years from 2005-2011 belong to the CENG, whereas only 20% of all students that graduated were from that college. There is a similar but opposite behavior for the students of the OCOB, 23.7% of all graduating students were a member of the business college, but of all those that had not graduated by Spring 2011, only 13.3% belonged to OCOB. We also see a less pronounced difference in the *Gender* variable. Males have a higher contribution to the total number of DropOut/Still Enrolled group (63.2% Male, 36.8% Female) and conversely a higher percentage of females graduated in the six year period (51.4% to 48.6%).

Turning our attention to the *Ethnicity*, *FirstGeneration*, *Geography*, and *Remed* variables, we see that the relative percentage contribution of each level to each category for both the Graduation and DropOut/Still Enrolled groups are approximately equal. This seems to indicate that a strong relationship between these background characteristics and graduation rates may not exist.

Now, let's compare the switching major behaviors for each of these groups. This information is provided in Table 4. Major switching rates and switching behaviors vary tremendously between the colleges. We see that although only about 21.2% of all students belong to the OCOB, they are responsible for 24.9% of all students that persist in their original major. Similarly, 23% of Cal Poly students entered in the CENG but they are responsible for 52.8% of all major switches that are to STEM majors. Finally, though only 16.3% and 19.1% of all students belonged to CLA and CAFES respectively, they both are responsible for about 30% of all major switches to non-STEM majors. Turning our attention to the *Gender* variable, we see that though almost equal percentages of males and females persist in their majors, the switching behaviors are very different, about 60% of all major switching by males is to a STEM major whereas about 60% of all major switching by females is to a non-STEM major.

With respect to the other variables, we see very comparable major switching rates and behaviors through each level.

| Variable | Levels | nPersist | %Persist | nSwSTEM | %SwSTEM | nSwNon-STEM | %SwNon-STEM | nall | %all |
|-----------------|-------------------|----------|----------|---------|---------|-------------|-------------|------|-------|
| College | CAED | 254 | 9.4 | 19 | 5.6 | 33 | 8.4 | 306 | 8.9 |
| | CAFES | 491 | 18.2 | 49 | 14.4 | 114 | 29.0 | 654 | 19.1 |
| | CENG | 562 | 20.9 | 179 | 52.8 | 47 | 12.0 | 788 | 23.0 |
| | CLA | 422 | 15.7 | 26 | 7.7 | 109 | 27.7 | 557 | 16.3 |
| | CSM | 294 | 10.9 | 47 | 13.9 | 53 | 13.5 | 394 | 11.5 |
| | OCOB | 670 | 24.9 | 19 | 5.6 | 37 | 9.4 | 726 | 21.2 |
| | all | 2693 | 100.0 | 339 | 100.0 | 393 | 100.0 | 3425 | 100.0 |
| Ethnicity | AtAm | 24 | 0.9 | 3 | 0.9 | 6 | 1.5 | 33 | 1.0 |
| | AsAm | 295 | 10.9 | 50 | 14.8 | 34 | 8.6 | 379 | 11.1 |
| | Haw | 12 | 0.4 | 2 | 0.6 | 1 | 0.2 | 15 | 0.4 |
| | Hisp | 259 | 9.6 | 30 | 8.8 | 30 | 7.6 | 319 | 9.3 |
| | Inter | 17 | 0.6 | 0 | 0.0 | 0 | 0.0 | 17 | 0.5 |
| | NaAm | 25 | 0.9 | 2 | 0.6 | 4 | 1.0 | 31 | 0.9 |
| | Unknown | 246 | 9.1 | 33 | 9.7 | 34 | 8.6 | 313 | 9.1 |
| | White | 1815 | 67.4 | 219 | 64.6 | 284 | 72.3 | 2318 | 67.7 |
| | all | 2693 | 100.0 | 339 | 100.0 | 393 | 100.0 | 3425 | 100.0 |
| Gender | Male | 1423 | 52.8 | 212 | 62.5 | 150 | 38.2 | 1785 | 52.1 |
| | Female | 1270 | 47.2 | 127 | 37.5 | 243 | 61.8 | 1640 | 47.9 |
| | all | 2693 | 100.0 | 339 | 100.0 | 393 | 100.0 | 3425 | 100.0 |
| FirstGeneration | FirstGen | 2448 | 90.9 | 313 | 92.3 | 357 | 90.8 | 3118 | 91.0 |
| | NotFirstGen | 245 | 9.1 | 26 | 7.7 | 36 | 9.2 | 307 | 9.0 |
| | all | 2693 | 100.0 | 339 | 100.0 | 393 | 100.0 | 3425 | 100.0 |
| Remed | AtLeastSomeREMEDI | 344 | 12.8 | 42 | 12.4 | 52 | 13.2 | 438 | 12.8 |
| | NoREMEDI | 2349 | 87.2 | 297 | 87.6 | 341 | 86.8 | 2987 | 87.2 |
| | all | 2693 | 100.0 | 339 | 100.0 | 393 | 100.0 | 3425 | 100.0 |

Table 4: Background Characteristics Broken Down by Major Switching Behavior for 2005 Cal Poly Freshman Class

Now, let's quickly look at the breakdown in the quantitative variables *High School GPA* and *SAT Score* by graduation and major switching behaviors. This information is displayed in Tables 5 & 6.

| Variable | Levels | \bar{x} | \tilde{x} | s |
|----------|------------------------|-----------|-------------|-------|
| HSGPA | Graduated | 3.8 | 3.8 | 0.4 |
| | DropOut/Still Enrolled | 3.6 | 3.6 | 0.4 |
| | all | 3.7 | 3.8 | 0.4 |
| SAT | Graduated | 1204.6 | 1210.0 | 119.6 |
| | DropOut/Still Enrolled | 1202.4 | 1210.0 | 137.3 |
| | all | 1204.0 | 1210.0 | 124.1 |

Table 5: Academic Background by Graduation Status

| Variable | Levels | \bar{x} | \tilde{x} | s |
|----------|-------------|-----------|-------------|-------|
| HSGPA | Persist | 3.7 | 3.8 | 0.4 |
| | Sw STEM | 3.8 | 3.8 | 0.4 |
| | Sw Non-STEM | 3.7 | 3.7 | 0.4 |
| | all | 3.7 | 3.8 | 0.4 |
| SAT | Persist | 1204.3 | 1210.0 | 124.8 |
| | Sw STEM | 1230.4 | 1240.0 | 123.5 |
| | Sw Non-STEM | 1179.6 | 1180.0 | 115.1 |
| | all | 1204.0 | 1210.0 | 124.1 |

Table 6: Academic Background by Major Switch

By examining the tables we see that there are is not any substantial differences in High school GPA or SAT scores when comparing switching behaviors, however, when looking at graduation there does seem to be a noticeable difference between those that graduated and those that had not after 6 years when looking at the *High School GPA* variable. Students that ultimately graduated from Cal Poly within the six year period have a slightly higher mean and a trimmed mean High School GPAs.

Analysis/Results

Multinomial Logistic Regression Model

Our first research goal was to identify background characteristics and academic performance measures that are associated with various switching behaviors. To aid us in this analysis we began by creating a variable that we named *FirstSwitch* that has three nominal levels. The levels corresponded to different possible switching behaviors that we were interested in studying. The first level represent those that persisted (never switched majors), the second level represent those whose first major switch was into a STEM major, and the final level represent those whose first major switch was to a non-STEM major. To successfully accomplish the goal of identifying characteristics that are associated with various major switching behaviors we need to find a way to relate this nominal categorical variable *FirstSwitch* to the categorical and continuous predictors. An effective statistical tool for analyzing relationships between a categorical nominal dependent variable and categorical or continuous independent variables is multinomial logistic regression, and we made use of it to accomplish the first part of our analysis.

Multinomial regression allows us to assess which independent variables are useful in determining the odds that a student will switch majors rather than persist but it also allows us to see how the relationships between various predictors change when estimating the odds the major switch would be to a STEM versus a non-STEM major. Multinomial regression output provides us two sets of coefficients, one set for each comparison that needs to be made. One set to compare those that switch to STEM majors to those that persist and another set to compare those that switch to a non-STEM major to those that persist. For this analysis, it makes the most sense to make those that persist in their original major the baseline group because it is the most natural to interpret. The multinomial regression model assumes that the natural logarithm of the odds of a student being in a major switching group to persisting is a linear function of the independent variables. In this case we have:

$$\ln\left(\frac{\Pr(\text{Switch to STEM})}{\Pr(\text{Persist})}\right) = \mathbf{B}_1^T * \mathbf{X} \text{ and } \ln\left(\frac{\Pr(\text{Switch to non-STEM})}{\Pr(\text{Persist})}\right) = \mathbf{B}_2^T * \mathbf{X},$$

where \mathbf{X} is the vector of predictors and the \mathbf{B}_i is the vector of parameters coefficients

Multinomial Regression is sensitive to having combinations of predictor variables with few observations. For this reason we had to make a few adjustments, both to our variables and our dataset. First, we removed *College* and used *STEM Matric* to indicate whether the student's initial major was STEM or non-STEM. The method of determining which majors were considered to be a STEM major was guided by the classifications made by Chen (2009) as well as a close inspection of the coursework required for each major using the Cal Poly Course Catalog. A table listing all Cal Poly majors and our corresponding classifications is located in Table 13 in the appendix on page 30. Secondly, we had to cut out groups of students that had small populations. This included those of Hawaiian/Pacific Islander, African American, and Native American ethnicity. Students with missing values in any predictor were also removed from the analysis.

The mlogit package was used to obtain maximum likelihood estimates of the model parameters. Results are available in Tables 7 below. To assess model fit, a global likelihood ratio test is carried out.

For this dataset, the resulting test statistic had a very small p-value and thus we can conclude that the model is useful in differentiating major switching behavior. Further, drop in deviance tests can be conducted for each predictor in the model to test whether or not the independent variables have an overall relationship to the response variable, *FirstSwitch*. The drop in deviance test statistic is computed by fitting two models: one model with all the relevant parameters and a second model that excludes the predictors being tested, thus the first model has g more parameters being estimated. The null hypothesis is that the parameter estimates corresponding to predictors included in the full model but not the reduced model are zero and the alternative hypothesis is that at least one of the parameter estimates is not zero. The test statistic is:

$$D = 2(\text{LogLikelihood}_{\text{Full Model}} - \text{LogLikelihood}_{\text{Reduced Model}})$$

Under the null hypothesis, D is approximately chi-square distributed with g degrees of freedom. Once it has been determined that a predictor has an overall relationship to the response then we can look at the significance of each predictor in being able to distinguish between those that switch majors to STEM and those that persist, as well as distinguishing those that switch to non-STEM majors and those that persist.

Multinomial Regression models carry with them the assumption of independence of irrelevant alternatives. Essentially this assumption asserts that a person's preference between two alternatives is not affected by the presence of another alternative. In the context of this project for example, the amount a student prefers to switch to a STEM major rather than to persist in their matriculating major is not dependent on the option to switch to a non-STEM major. The Hausman-McFadden test was carried out to test this assumption. The resulting p-value of the test was very near one, we have no reason to believe that this assumption has been violated in this case.

The Discrete Time Hazard Model

Of the four main research questions previously laid out, there are still three that we have yet to answer. Instead of trying to identify who might be likely to change majors, the remaining three questions were created to guide us in exploring the relationship between switching majors and the risk of graduation from Cal Poly. To be able to answer research questions (2) – (4) we need to be able to take the following circumstances into consideration.

Since the data covers the incoming 2005 Cal Poly freshman class only over a six year period of time we have no way of knowing if students that have not graduated by summer quarter 2011 will eventually do so. This introduces the problem of censoring. Ignoring censored observations has very dire analytical consequences and a branch of statistics has been developed to deal with censoring, survival analysis. Most survival analysis techniques, however, are designed to be used on continuous time to event random variables. In this study the event of interest, graduating from Cal Poly, can only happen at the end of a quarter. This makes the time to event random variable for this analysis discrete.

In addition, the data we have been provided contains both time varying and time invariant predictors. A time invariant predictor is one that remains constant over time, such as *Gender*. A time varying predictor is one that changes value over time. The variable *Units* is an example of a time varying predictor as the unit load a student takes can vary from quarter to quarter, though note, the values of all predictors, even time varying predictors, are frozen during each quarter. We want to be able to make use of both types of predictors in our analysis.

Finally, we note that graduation is a non-repeatable event. Once a student graduates, they are ineligible to graduate at a later date. This makes time to graduation inherently conditional in the sense that you are only eligible to graduate in quarter j given that you have not already graduated in quarters $1, 2, \dots, j-1$.

A model introduced by the famous statistician Dr. David Cox in the 1970's and elucidated by many researchers over the years, most notably Dr. Judith Singer and Dr. John Willett of Harvard, can handle all of these problems simultaneously. This model is known as a Discrete-Time Hazard Model (DTHM). Hazard, a common quantity in survival analysis, is the “backbone” of the DTHM model. The benefit to using the hazard function is that it allows us to assess the risk of graduation in every quarter. In this case, we define hazard to be the conditional probability that a student graduates in some quarter j given that the student did not graduate in the previous $j-1$ quarters. (Note, in traditional survival analysis with a continuous time to event random variable hazard is a rate but in a DTHM hazard is a probability.) If we define random variable T to indicate the time period j when a randomly selected student graduates, then we can express hazard as:

$$h_j = \Pr(T = j | T \geq j)$$

We are interested in identifying predictors that help to determine whether different types of students have different hazard functions. To do this we must include the various predictors into our definition of hazard. The predictor adjusted hazard can be defined as the conditional probability that

student i , $i=1,2,\dots,N$ (N = number of student in sample), graduates in quarter j , $j=1,2,\dots,J$ (J = maximum number of quarters a student can be in the dataset, 24 quarters), given values for each of P predictor values $z_{ij} = (z_{1ij}, \dots, z_{Pij})$

$$h_{ij} = \Pr(T_i = j | T_i \geq j, Z_{1ij} = z_{1ij}, \dots, Z_{Pij} = z_{Pij})$$

The most general form of the DTHM assumes that logistic transformation of hazard is a linear function of the P predictors and an intercept (represented by the α 's with use of time dummies D) where α_j is the intercept for quarter j . We can write this formally as:

$$\ln\left(\frac{h_{ij}}{1 - h_{ij}}\right) = (\alpha_1 D_{1ij} + \dots + \alpha_{24} D_{24ij}) + (\beta_1 Z_{1ij} + \dots + \beta_P Z_{Pij})$$

Where $D_{hij} = 1$, if $h = j$ and zero otherwise.

By solving for h_{ij} we obtain:

$$h_{ij} = \frac{1}{1 + e^{-[(\alpha_1 D_{1ij} + \dots + \alpha_{24} D_{24ij}) + (\beta_1 Z_{1ij} + \dots + \beta_P Z_{Pij})]}}$$

To obtain parameter estimates we turn to the method maximum likelihood. As usual in survival analysis, the likelihood function is the product of the probabilities of observing the sample data. If we introduce a vector called the event indicator, y_{ij} where for student i , $y_{ij} = 1$ if the student graduates in time period j and is zero otherwise, and let j_i represent the total number of quarters that have passed since Fall 2005 in which student I had yet to graduate Cal Poly, then it can be shown that the likelihood function is:

$$L = \prod_{i=1}^n \prod_{j=1}^{j_i} h_{ij}^{y_{ij}} (1 - h_{ij})^{(1-y_{ij})}$$

Here we see that for each time period, a student either graduates or does not. If they do graduate, then at time j_i then they contribute h_{ij} to the likelihood function, otherwise they contribute $(1-h_{ij})$. This likelihood function is identical to a sequence of independent Bernoulli trials in parameters h_{ij} , where the number of trials is $K = j_1 + j_2 + \dots + j_n$, the sum across all students, of the quarters each student is at Cal Poly without graduating. This equivalence allows for easy maximum likelihood estimation of the model parameters using standard logistic regression routines. The only caveat is that the data set needs to be transformed into what is known as “person period” form. A more detailed explanation of this and an example is located in Figure 3 of the appendix on page 29.

After parameter estimates have been obtained we need to assess model fit. Traditionally, the first model fit in a DTHM is the baseline hazard, the model that results from only using time as a predictor. If additional predictors are then added to the model we can assess if their inclusion into the model helps to fit the data better. In logistic regression we can accomplish this using a drop in deviance test, as described earlier.

In survival analysis, it is usually of interest to obtain an estimate of the survival function. To “survive” in this context means to “not graduate.” Once hazard estimates are obtained, the corresponding estimated survival probabilities at time j can be easily calculated as:

$$\hat{S}_j = \prod_{k=1}^j (1 - \hat{h}_k)$$

The mean time to graduation in can be easily estimated by: $\widehat{E(T)} = \sum_{i=1}^{24} \hat{S}_i$

For this part of the analysis we worked with an initial sample size of 46,412 person period records for 3012 students. Of the original 3425 students, we excluded students that declared themselves to be international or unknown in the *Ethnicity* variable, and students that had missing values in any of the predictors.

In order to have a more parsimonious model, a quartic polynomial model in time was fit. As is traditional, the first model fit was a model that only included time, baseline hazard. After estimating the baseline hazard we introduced the demographic background characteristics into the model. A drop in deviance test showed that the demographic and academic background characteristics have helped provide a better fit to the data. Next, information about Cal Poly academic performance and behavior, were added to the model and again a drop in deviance test identified these predictors as helping to improve model fit. Finally, interactions between *FirstSwitch.Vary* and *Ethnicity*, *Year Switch*, *Gender*, and *College* were added as well as interactions between time and GPA and time and Units. Again, these were found to improve model fit. Results of the test are located in Table 9. Now that we have a useful model we can take a look at the research questions and see what the model says.

Table 9: Discrete Time Hazard Model Results

| | Model A | Model B | Model C | Model D |
|---|----------------|----------------|----------------|----------------|
| Model df | 5 | 16 | 22 | 67 |
| AIC | 11341 | 11079 | 10438 | 9937.2 |
| LL | -5665.4 | -5523.6 | -5196.8 | -4901.6 |
| -2(Δ LL) | - | 283.55 | 653.73 | 590.29 |
| df | - | 11 | 6 | 45 |
| P-Value Comparing A to B B to C C to D | - | <.001 | <.001 | <.001 |

Model A: Only time as a predictor

Model B: Time and Background (Academic and Demographic)

Model C: Time, Background (Academic and Demographic), and Academic Performance at Cal Poly

Model D: Time, Background (Academic and Demographic), Academic Performance at Cal Poly, and relevant interactions.

Discussion

Interpretation of the Multinomial Results

The first research question was designed to assess background characteristics associated with various major switching behaviors. The predictors that are useful both overall and to distinguish those switching to STEM majors and those that persist are *STEM Matric* and *Fall05 GPA*. Students that enter Cal Poly in a STEM major are more likely to switch to another stem major than to persist in their original majors. Further, the higher a student's *Fall 2005 GPA* the more likely that student is to switch to a STEM major rather than persist. The *Geography* variable is just barely not significant at the $\alpha = .05$ significance level but it suggests that California residents are slightly more likely to switch to a STEM major rather than persist.

The predictors that are useful both overall and to distinguish those switching to non-STEM majors and those that persist are *STEM Matric*, *Gender*, *High School GPA*, *SAT score*, *Fall05 GPA*, and the interaction term *STEM Matric*Fall05 GPA*. Those that enter as a STEM major are more likely to switch to a non-STEM major than to persist. Males are less likely to switch to a non-STEM major rather than persist. The higher a student's High School GPA the less likely they are to switch to a non-STEM major rather than persist, however, the higher a student's Fall 2005 GPA the more likely that student is to switch to a non-STEM major rather than persist, however, this effect is stronger for those entering Cal Poly in STEM majors when compared to those that entered in non-STEM majors.

Table 7: Drop in Deviance Tests of Significance for Predictors Included in the Multinomial Logistic Regression

| Predictor | Log Likelihood | Drop in Deviance χ^2 | DF | P-Value |
|-----------------------------|----------------|---------------------------|----|---------|
| STEM Matric | -1900.8 | 138.17 | 6 | <.001 |
| Gender | -1842.1 | 20.90 | 2 | <.001 |
| Ethnicity | -1835.9 | 8.40 | 4 | 0.08 |
| First Generation | -1831.9 | 0.56 | 2 | 0.76 |
| High School GPA | -1838.7 | 14.02 | 4 | 0.007 |
| SAT Score | -1836.2 | 9.030 | 2 | 0.011 |
| California Residency | -1834.6 | 5.96 | 2 | 0.051 |
| Remedial Work | -1832.7 | 2.14 | 2 | 0.34 |
| Fall 05' Units | -1831.9 | 0.47 | 2 | 0.79 |
| Fall 05' GPA | -1852.6 | 20.90 | 4 | <.001 |
| Pell Grant Recipient | -1831.9 | 0.43 | 2 | 0.81 |
| STEM:High School GPA | -1832 | 0.58 | 2 | 0.75 |
| STEM:Fall 05 GPA | -1838.4 | 13.42 | 2 | <.001 |
| Full Model | -1831.7 | 238.42 | | <.001 |

Table 8: Multinomial Model Results

| Predictor | Estimate | Std. Error | Odds Ratio | P-Value |
|--|----------|------------|------------|---------|
| <u>Compare Persist to Switch STEM</u> | | | | |
| 1:(intercept) | -5.62 | 1.63 | 0.00 | <.001 |
| 1:STEM Matric | 3.82 | 1.62 | 45.72 | 0.018 |
| 1:Male | 0.18 | 0.14 | 1.20 | 0.19 |
| 1:Hispanic/Latino | -0.32 | 0.26 | 0.73 | 0.22 |
| 1:White | -0.33 | 0.18 | 0.72 | 0.06 |
| 1:First Generation | -0.12 | 0.28 | 0.88 | 0.66 |
| 1:High School GPA | 0.087 | 0.43 | 1.09 | 0.84 |
| 1:SAT Score | 0.000033 | 0.00066 | 1.00 | 0.95 |
| 1:Out Of State | 0.40 | 0.19 | 1.49 | 0.04 |
| 1:No Remedial Work | -0.077 | 0.23 | 0.93 | 0.74 |
| 1:Fall05 Units | 0.027 | 0.04 | 1.03 | 0.50 |
| 1:Fall05 GPA | 0.70 | 0.25 | 2.02 | 0.0055 |
| 1:Received Pell Grant | 0.1447 | 0.22 | 1.16 | 0.51 |
| 1:STEM Matric*High School GPA | -0.26 | 0.47 | 0.77 | 0.58 |
| 1:STEM Matric*Fall05 GPA | -0.44 | 0.27 | 0.64 | 0.11 |
| <u>Compare Persist to Switch non-STEM</u> | | | | |
| 2:(intercept) | -0.038 | 1.11 | 0.96 | 0.97 |
| 2:STEM Matric | 2.93 | 1.18 | 18.80 | 0.013 |
| 2:Male | -0.53 | 0.13 | 0.59 | <.001 |
| 2:Hispanic/Latino | -0.57 | 0.27 | 0.94 | 0.83 |
| 2:White | 0.29 | 0.19 | 1.34 | 0.14 |
| 2:First Generation | -0.183 | 0.28 | 0.83 | 0.52 |
| 2:High School GPA | -0.56 | 0.26 | 0.57 | 0.03 |
| 2:SAT Score | -0.0019 | 0.00063 | 1.00 | 0.003 |
| 2:Out Of State | -0.27 | 0.24 | 0.76 | 0.27 |
| 2:No Remedial Work | 0.28 | 0.21 | 1.33 | 0.17 |
| 2:Fall05 Units | -0.0022 | 0.036 | 0.99 | 0.95 |
| 2:Fall05 GPA | 0.73 | 0.15 | 2.08 | <.001 |
| 2:Received Pell Grant | 0.018 | 0.22 | 1.02 | 0.93 |
| 2:STEM Matric*High School GPA | -0.20 | 0.35 | 0.82 | 0.56 |
| 2:STEM Matric*Fall05 GPA | -0.66 | 0.20 | 0.52 | <.001 |

Interpretation of DTHM Results

The second and third main research questions sought to determine if switching major is associated with the risk of graduating. Predictors in the model associated with major switching behavior are statistically significant and these include the interactions between *FirstSwitch.vary* and *YearSwitch*, *Gender*, and *College*. Thus we have evidence that the association between changing major and the risk of graduation is complicated and is dependent on the student's matriculating college, gender, the year in which the switch was made, and whether the first switch was made to a STEM major or a non-STEM major. Though this is a partially satisfying answer, it turns out that we can make some more insightful conclusions. If we were to look through all the majors in each college at Cal Poly, we could roughly separate the colleges into "Science" and "non-Science" colleges. The "science" colleges would be CAFES, CSM, CAED, and CENG leaving the "non-science" colleges CLA and OCOB. For the "science" colleges it turns out that if the student is going to change their major in the first three years¹, the hazard of graduation generally increases with a major change, the increase in hazard is always highest when switching to non-STEM majors. For the "non-STEM" colleges, persisting in the original major has the highest hazard of graduation for these colleges, followed by major switches to non-STEM majors.

To obtain a visual representation of this we will introduce a "typical" student profile and allow other predictors to vary. For this purpose we will fix the various background predictors at the following levels unless otherwise stated:

- *Gender*: male
- *Geography*: California resident
- *Pell Grant Status*: does not receive Pell grants
- *Ethnicity*: White
- *FirstGeneration*: not a first generation student
- *High School GPA*: 3.75(median value)
- *SAT Score*:1200 (median value)
- *Units*: 12 per quarter
- *GPA*: 3.0
- *Switch Yr*: 3(Median Switch Year)

Allowing both *FirstSwitch.vary* and *College* to vary and holding all the other predictors fixed to the values described above, we can obtain a nice picture of the hazard estimates plotted by quarter.

Taking a look at Figure 1a, it is clear to see that switching majors greatly increases the estimated hazard of graduation for students that matriculated to "science" colleges. For CENG, this jump in hazard of graduation is dramatic when comparing those that persist to those that switched to non-STEM in the third year. We see that for OCOB and CLA the reverse behavior holds, switching to a STEM major in the third year approximately halves the hazard of graduation. This fits well with observations made in the descriptive statistics section. These patterns hold for switches made in years one and two, as well.

¹ Most majors at Cal Poly are designed to be completed in four years. It seems logical that switching majors in the fourth or fifth years would not result in higher hazard graduation.

Table 10: DTHM Results

| Variable | Coef (Logit Scale) | Exp(Coef) | Individual Significance at $\alpha = .05$ |
|--------------------------------|---------------------------|------------------|---|
| Intercept | -22.523 | <.001 | Yes |
| Time | -4.841 | 0.008 | |
| Time Squared | 0.867 | 2.38 | |
| Time Cubed | -0.046 | 0.955 | |
| Time Fourth | 0.001 | 1.001 | |
| | | | |
| First Generation Student | -0.207 | 0.813 | No |
| Not a California Resident | -0.147 | 0.864 | No |
| Receive Pell Grant | -0.235 | 0.79 | Yes |
| Male | -0.418 | 0.658 | Yes |
| | | | |
| Asian-American | -0.08 | 0.923 | No |
| Hawaiian/Pacific Islander | -0.44 | 0.644 | |
| Hispanic/Latino | 0.048 | 1.049 | |
| Native American | 0.41 | 1.506 | |
| White | -0.029 | 0.972 | |
| | | | |
| | | | |
| HSGPA | 0.568 | 1.765 | Yes |
| SAT Combined Score | <.001 | >.999 | No |
| No Remedial Work Required | 0.107 | 1.113 | No |
| | | | |
| Number of Major Changes | -0.668 | 0.512 | Yes |
| First Switch to STEM Major | 7.558 | 1915.767 | Yes |
| First Switch to non-STEM Major | 9.296 | 10897.181 | |
| Units Attempted | 0.07 | 1.072 | Yes |
| GPA | 1.396 | 4.04 | Yes |
| | | | |
| CAFES | 0.897 | 2.453 | |
| CENG | 0.311 | 1.365 | |
| CLA | 1.547 | 4.697 | Yes |
| CSM | 0.912 | 2.49 | |
| OCOB | 1.853 | 6.377 | |
| | | | |
| | | | |
| T:GpaVector | -0.042 | 0.959 | Yes |
| T:Units | -0.009 | 0.991 | Yes |
| | | | |

| | | | |
|---|--------|----------|-----|
| CAFES: Switch to STEM | -0.033 | 0.968 | Yes |
| CENG:Switch to STEM | -0.54 | 0.583 | |
| CLA:Switch to STEM | -1.13 | 0.323 | |
| CSM:Switch to STEM | 0.048 | 1.049 | |
| OCOB: Switch to STEM | -1.515 | 0.22 | |
| CAFES: Switch to non-STEM | 0.108 | 1.114 | |
| CENG:Switch to non-STEM | 0.387 | 1.473 | |
| CLA:Switch to non-STEM | -0.607 | 0.545 | |
| CSM:Switch to non-STEM | -0.106 | 0.899 | |
| OCOB: Switch to non-STEM | -0.928 | 0.395 | |
| | | | |
| Asian-American:Switch to STEM | 1.639 | 5.149 | No |
| Asian-American:Switch to non-STEM | -0.791 | 0.453 | |
| Hawaiian/Pac Islander: Switch to STEM | 2.132 | 8.432 | |
| Hawaiian/Pac Islander: Switch to non-STEM | 1.184 | 3.267 | |
| Hispanic/Latino: Switch to STEM | 0.726 | 2.066 | |
| Hispanic/Latino: Switch to non-STEM | -0.638 | 0.528 | |
| Native American: Switch to STEM | 0.604 | 1.83 | |
| Native American: Switch to non-STEM | -1.14 | 0.32 | |
| White: Switch to STEM | 1.298 | 3.663 | |
| White: Switch to non-STEM | -0.758 | 0.468 | |
| | | | |
| Male: Switch to STEM | 0.241 | 1.272 | Yes |
| Male: Switch to non-STEM | 0.101 | 1.107 | |
| | | | |
| Persist:Switch Yr 2 | 0.769 | 2.158 | Yes |
| Switch to STEM: Switch Yr 2 | -0.291 | 0.748 | |
| Switch to non-STEM: Switch Year 2 | 0.411 | 1.509 | |
| Persist:Switch Yr 3 | 0.563 | 1.756 | |
| Switch to STEM: Switch Yr 3 | -0.623 | 0.536 | |
| Switch to non-STEM: Switch Year 3 | 0.045 | 1.046 | |
| Persist:Switch Yr 4 | -4.469 | 0.011 | |
| Switch to STEM: Switch Yr 4 | -0.917 | 0.4 | |
| Switch to non-STEM: Switch Year 4 | -0.594 | 0.552 | |
| Persist:Switch Yr 5 | -7.276 | 0.001 | |
| Switch to STEM: Switch Yr 5 | -1.261 | 0.283 | |
| Switch to non-STEM: Switch Year 5 | -0.559 | 0.572 | |
| Persist:Switch Yr 6 | -8.694 | <0.001 | |
| Switch to STEM: Switch Yr 6 | -0.361 | 0.697 | |
| Switch to non-STEM: Switch Year 6 | -20.46 | <0.001 | |
| Persist:No Switch | 7.711 | 2232.502 | |

Also, for CENG we notice that if the major switch was made to a STEM major the hazard of graduation actually drops substantially. This actually makes sense for a few reasons. First, engineering majors tend to take students longer to complete, five years is a very typical. In addition, engineering students have a tendency to switch to other majors within the engineering college. So, since all engineering majors fall under the STEM umbrella then it seems logical that we would see switches to STEM majors resulting in lower hazard graduation.

Taking a look at Figure 1b, we see this same phenomenon displayed a little differently. We see that for the “non-science” colleges, persisting in ones matriculating major has between two and three times higher hazard graduation. Another interesting thing that we can see in Figure 1b is that there are three grouping of colleges that have similar hazard estimates in all three cases: OCOB and CLA, CENG and CAED, and CSM and CAFES.

Another perspective that was to be addressed by the last research question was to investigate how the timing of the first major switch affects the risk of graduation. A visual representation of this is available in Figure 2a, as well as a table of estimated mean survival times in Table 11.

Again, there are two distinct behaviors that are noticeable, one for the “science” and one for the “non-science” colleges. Switching in the first three years is associated with higher risk of graduation for those matriculating into “science” colleges. Further, switching to a non-STEM major has the highest risk of graduation if done in the second year where as for those switching to STEM majors the highest risk of graduation is obtained if done in the first year. It is true that having a higher risk of graduating is equivalent to having smaller estimated mean time to graduation and this is demonstrated in Table 11.

The behavior for the “non-science” colleges previously mentioned holds as well. Students who persist tend to graduate sooner than those that switch, regardless of STEM classification. An interesting thing to note however is that the optimal year for switching by type of major change is the same for the “science” and “non-science” colleges, hazard of graduation is highest when switching to STEM majors in the first year, and highest when switching to non-STEM in the second year.

Hazard by College by Switching Status

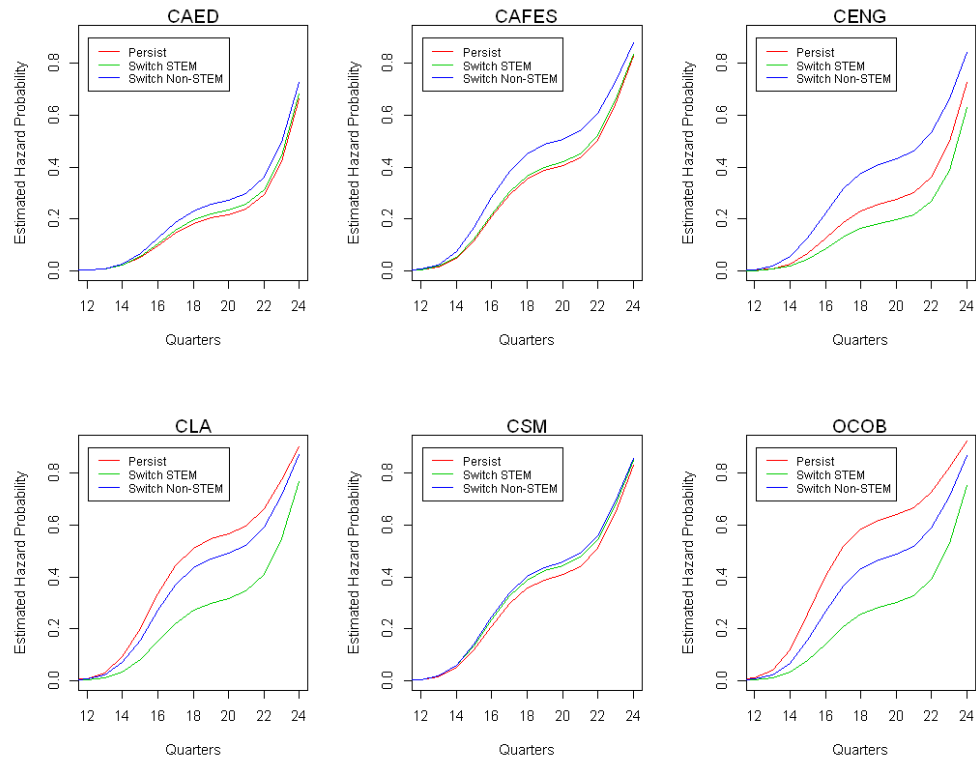


Figure 1a: Comparison of major switching behaviors by matriculating college.

Hazard by College by Switching Status

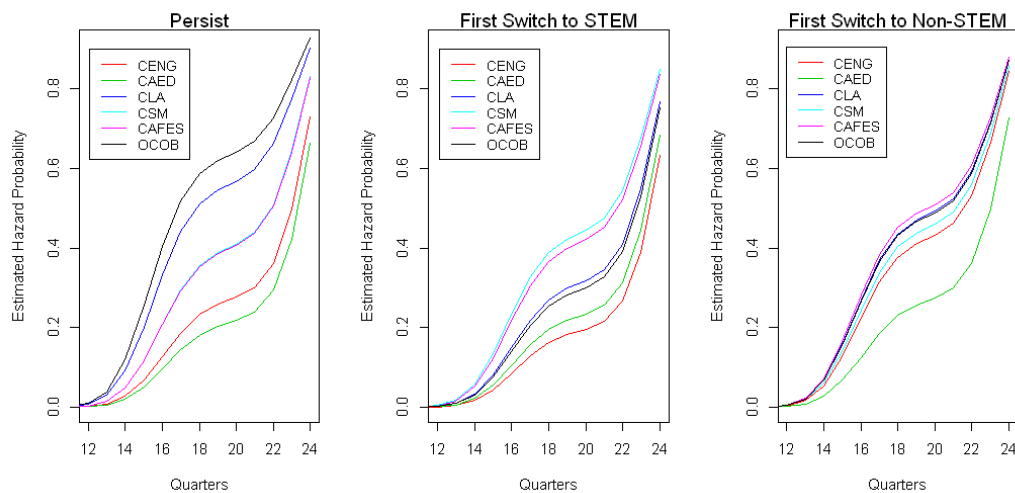


Figure 1b: Comparison of matriculating college by major switching behavior

| (SW Yr 1) | Persist | Switch STEM | Switch non-STEM |
|-----------|---------|-------------|-----------------|
| CENG | 17.75 | 17.34 | 16.42 |
| CAED | 18.45 | 16.88 | 17.86 |
| CLA | 15.46 | 16.11 | 16.01 |
| CSM | 16.51 | 15.28 | 16.23 |
| CAFES | 16.53 | 15.41 | 15.90 |
| OCOB | 15.06 | 16.24 | 16.03 |
| (SW Yr 2) | | | |
| CENG | 17.75 | 18.00 | 15.74 |
| CAED | 18.45 | 17.48 | 16.97 |
| CLA | 15.46 | 16.63 | 15.38 |
| CSM | 16.51 | 15.71 | 15.57 |
| CAFES | 16.53 | 15.86 | 15.29 |
| OCOB | 15.06 | 16.78 | 15.39 |
| (SW Yr 3) | | | |
| CENG | 17.75 | 18.78 | 16.34 |
| CAED | 18.45 | 18.23 | 17.76 |
| CLA | 15.46 | 17.30 | 15.93 |
| CSM | 16.51 | 16.25 | 16.15 |
| CAFES | 16.53 | 16.42 | 15.83 |
| OCOB | 15.06 | 17.46 | 15.96 |
| (SW Yr 4) | | | |
| CENG | 17.75 | 19.48 | 17.63 |
| CAED | 18.45 | 18.94 | 19.26 |
| CLA | 15.46 | 17.95 | 17.10 |
| CSM | 16.51 | 16.79 | 17.38 |
| CAFES | 16.53 | 16.98 | 16.97 |
| OCOB | 15.06 | 18.13 | 17.13 |

Table 11: Estimated Mean Time to Graduation (In Quarters)

Though the main research questions have been more or less answered, it is time to take a look at a few other interesting findings that our model helps to reveal. The first thing of interest is that women have a higher hazard of graduation than men and thus lower estimated time until graduation, regardless of major switch behavior. This can be seen below in Table 12. Since differing behavior between “science” and “non-science” colleges has been so consistent, a representative college of each type was selected for comparison. The significant interaction between *Gender* and *FirstSwitch.vary* is also interesting in that shows that though women tend to graduate faster than men, switching majors regardless of STEM classification actually raises their expected mean time to graduation. In contrast, a major switch of either type is associated with a decrease in estimated mean time to graduation for males, but switching to a STEM major is associated with a larger decrease than switching to a non-STEM major.

Baseline Hazard by College by Year Switch, FS = 1

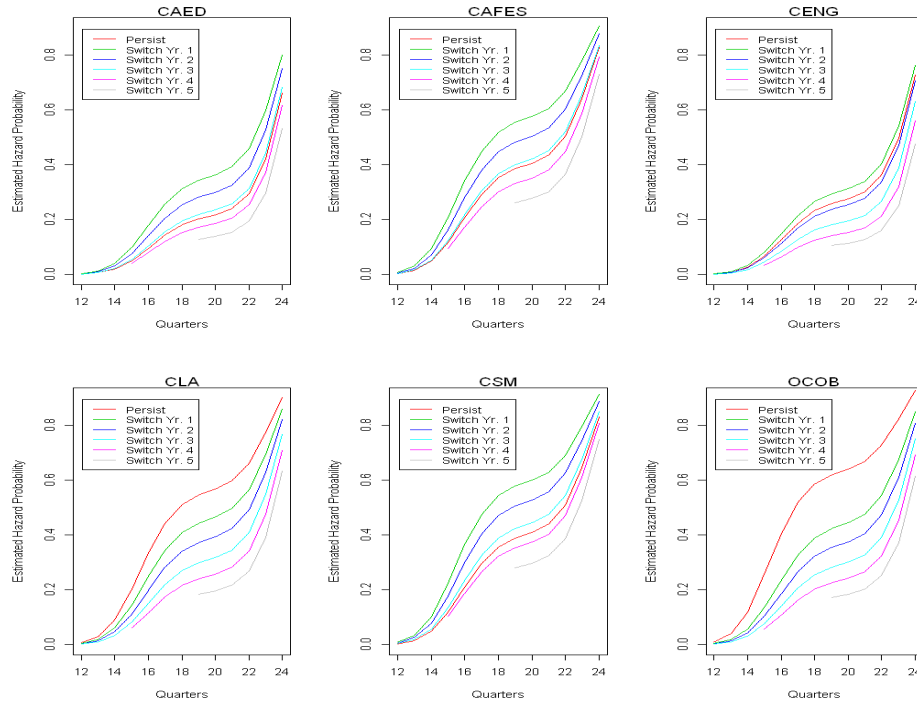


Figure 2a: Comparison of how timing of switching to STEM majors relates to hazard estimates

Baseline Hazard by College by Year Switch, FS = 2

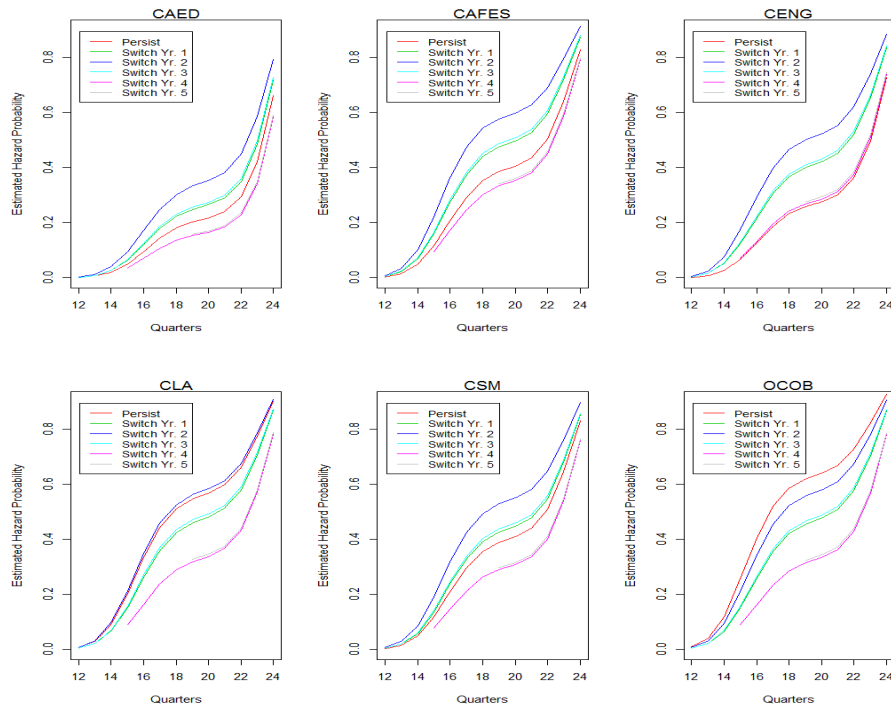


Figure 2b : Comparison of how timing of switching to non-STEM majors relates to hazard estimates

| Gender/College Type | Persist | Switch STEM | Switch non-STEM |
|---------------------|---------|-------------|-----------------|
| Male/SCI | 17.75 | 19.48 | 17.63 |
| Female/SCI | 16.86 | 19.07 | 16.95 |
| Male/Non-SCI | 15.46 | 17.95 | 17.10 |
| Female/Non-SCI | 14.91 | 17.55 | 16.49 |

Table 12: Estimated mean time to graduation, in quarters. SCI = CENG, Non-SCI = CLA

Another interesting thing to look at is what specific predictors were significantly associated with hazard of graduation individually, in the presence of the other predictors. Though they were included as control variables, it is still important to compare to the findings of other researchers. Interactions between Time and both *GPA* and *Units* were statistically significant. Higher GPA and higher unit loads are associated with higher hazard graduation initially but the interaction tells us that the association decreases with time. Also, students receiving Pell Grants were estimated at having a decrease in hazard of graduation. *High School GPA* was significantly associated with hazard of graduation but neither *SAT Score* nor *Remed* were. Higher High School GPAs are associated with much higher hazard of graduation. *Ethnicity*, *FirstGeneration*, and *Geography* were all found to be not significantly associated with hazard of graduation in the presence of the other predictors.

Limitations of Analyses

This study was conducted using data from a moderately selective, technically focused, public university in a relatively ethnically homogenous area. The results found may not generalize to other types of universities, however, the methodology used is flexible, allowing for similar analyses to be carried out with different datasets easily. The variables used are such that they should be readily available at most universities, allowing for investigation of how the association between major switching behaviors and the hazard of graduation might differ by characteristics of the university and student population.

There are some complications with the data that would have been nice to be able to circumvent. One big issue results from data not available to us. The dataset provided to us only had records of a student's primary major. Information about double majors and minors were not available. It seems reasonable that the inclusion of this information would allow a clearer picture of the relationship between hazard of graduation and the predictors used in the model. Unfortunately, there is no way of estimating the number of students that double major or minor making estimation of the size of the problem difficult. Another possible bit of information that could have been useful in our analysis regards student participation in study abroad programs or in Cal Poly athletics programs. Other types of analyses could have taken into consideration the effect that summer school attendance has on the risk of graduation.

Conclusion

The explicit purpose of this analysis was to illuminate the relationship between switching majors and degree attainment, with a particular emphasis on STEM majors. Data from the 2005 Cal Poly freshman cohort was used to assess this relationship while controlling for various background characteristics. Four research questions were developed in order to guide the analysis. These questions try to help identify what characteristics are associated with various major switching behaviors as well as assess how switching majors is related to college degree attainment while controlling for background characteristics, and how this relationship changes with the timing of the major change.

Extensive literature searches and review as well as a careful examination of descriptive statistics were performed to guide our expectations. The descriptive statistics showed what seemed to be different graduation rates for those switching majors and those that persisted in their original major. There were several other background characteristics that seemed to be related to degree attainment, most notably *College* and *Gender*.

In order to identify which background characteristics are associated with various major switching behaviors, a multinomial logistic regression model was created. The model is useful in determining how various background characteristics help to differentiate those that switch to STEM majors rather than persist as well as those that switch to non-STEM majors rather than persist. It was determined that students matriculating into a STEM major are more likely to switch majors than students matriculating to non-STEM majors, however, they are more likely to switch to another STEM major than to a non-STEM major. Further, it was shown that higher *Fall05 GPA* is associated with a higher chance of switching majors regardless of the STEM classification of the matriculating major. It was found that males are less likely than women to switch to a non-STEM major rather than persist. Higher *High School GPA* was found to lower the odds that a student would switch to a non-STEM major.

To assess the relationship between various major switching behaviors and degree attainment, a discrete time hazard model was fit. The DTHM allows for assessment of the hazard of graduation in every quarter while controlling for various background and academic performance measures. A useful model was constructed and results obtained. The relationship between hazard of graduation and major switching behavior was found to be complicated and depended on the student's matriculating college, gender, and the year that the switch occurred. This relationship between hazard of graduation and major switching behavior can be broken down into two basic varieties that depend on whether or not the student's matriculating college is a "science" or a "non-science" college. If a switch is made in the first three years, students matriculating to "science" colleges seem to benefit from switching majors in that it generally increases their hazard of graduation whereas persisting in one's original major is best for students matriculating to non-science majors. Another interesting finding was that if a student switches major, regardless of matriculating to a "science" college or not, hazard of graduation is increased the most if the switch takes place in the first year for switches to STEM majors and the second year for switches to non-STEM majors. Additionally, it was determined that women have higher hazard graduation regardless of major switching behavior than men but switching majors usually increases a women's time to graduation while decreasing it for men.

Appendix

Figure 3: Example Showing Conversion of Data from Person Form to Person Period Form

| In Person Form | | | | | | | | | | | | |
|-----------------------|---------|---------|--|-----|--------|--------|--------|--------|--------|--------|--------|------------|
| Student | Grad? | Censor? | Total Quarters | M/F | GPA Q1 | GPA Q2 | GPA Q3 | GPA Q4 | GPA Q5 | GPA Q6 | GPA Q7 | GPA Q8 |
| 1 | Yes | No | 6 | M | 3.45 | 3.6 | 3.5 | 3.6 | 2.9 | 3.6 | - | - |
| 2 | No | Yes | 2 | F | 3.5 | 2.9 | - | - | - | - | - | - |
| 3 | No | Yes | 8 | M | 3.3 | 2.75 | 3.56 | 3.61 | 3.15 | 2.85 | 3.2 | 3.5 |
| In Person Period Form | | | | | | | | | | | | |
| Quarter | Student | Y | M/F | D1 | D2 | D3 | D4 | D5 | D6 | D7 | D8 | GPA Vector |
| 1 | 1 | 0 | M | 1 | 0 | 0 | 0 | 0 | 0 | 0 | 0 | 3.45 |
| 2 | 1 | 0 | M | 0 | 1 | 0 | 0 | 0 | 0 | 0 | 0 | 3.6 |
| 3 | 1 | 0 | M | 0 | 0 | 1 | 0 | 0 | 0 | 0 | 0 | 3.5 |
| 4 | 1 | 0 | M | 0 | 0 | 0 | 1 | 0 | 0 | 0 | 0 | 3.6 |
| 5 | 1 | 0 | M | 0 | 0 | 0 | 0 | 1 | 0 | 0 | 0 | 2.9 |
| 6 | 1 | 1 | M | 0 | 0 | 0 | 0 | 0 | 1 | 0 | 0 | 3.6 |
| 1 | 2 | 0 | F | 1 | 0 | 0 | 0 | 0 | 0 | 0 | 0 | 3.5 |
| 2 | 2 | 0 | F | 0 | 1 | 0 | 0 | 0 | 0 | 0 | 0 | 2.9 |
| 1 | 3 | 0 | M | 1 | 0 | 0 | 0 | 0 | 0 | 0 | 0 | 3.3 |
| 2 | 3 | 0 | M | 0 | 1 | 0 | 0 | 0 | 0 | 0 | 0 | 2.75 |
| 3 | 3 | 0 | M | 0 | 0 | 1 | 0 | 0 | 0 | 0 | 0 | 3.56 |
| 4 | 3 | 0 | M | 0 | 0 | 0 | 1 | 0 | 0 | 0 | 0 | 3.61 |
| 5 | 3 | 0 | M | 0 | 0 | 0 | 0 | 1 | 0 | 0 | 0 | 3.15 |
| 6 | 3 | 0 | M | 0 | 0 | 0 | 0 | 0 | 1 | 0 | 0 | 2.85 |
| 7 | 3 | 0 | M | 0 | 0 | 0 | 0 | 0 | 0 | 1 | 0 | 3.2 |
| 8 | 3 | 0 | M | 0 | 0 | 0 | 0 | 0 | 0 | 0 | 1 | 3.5 |
| | | - | Gender is an example of a time invariant predictor. Since its value is always constant, then its value is repeated in the person period dataset for each quarter attended by the student. | | | | | | | | | |
| | | - | Term GPA is an example of a time varying predictor. The student's GPA for quarter j is listed in row j of the person period dataset. | | | | | | | | | |
| | | - | In the person form, all information about each student is located in a single row. In particular, there is information indicating whether or not a student ever graduates and their total number of quarters. This information can be turned into the Y vector that indicates for each quarter whether the student graduated or not. | | | | | | | | | |

Suppose that graduation at some university typically takes place in six quarters but information about eight quarters is known. We see that student 1 graduates in six quarters. This results in Student 1's record to be transformed into six rows of the person period dataset. Student 1's gender, a time invariant predictor, is repeated in each quarter. GPA, a time varying predictor, has the quarter j's GPA located in row j of the Student 1's person period record. The censor indicator and information about the total number of quarters attended by the student allow for creation of the Y vector. Students 2 and 3 are both censored; they have not graduated by the end of 8 quarters, the length of the data collection process. Student 2 only attended for two quarters (perhaps they dropped out) and student 3 was still taking courses at the end of the data collection process.

Table 13: Cal Poly Majors STEM Classification

| Major Abbreviation | Major Name | STEM Classification |
|---------------------------|--|----------------------------|
| AERO | Aerospace Engineering | STEM |
| AGB | Agribusiness | non-STEM |
| AGSC | Agricultural Science | STEM |
| ARCE | Architectural Engineering | STEM |
| ARCH | Architecture | non-STEM |
| ART | Art and Design | non-STEM |
| ASCI | Animal Science | STEM |
| ASM | Agricultural Systems Management | non-STEM |
| BCHM | Biochemistry | STEM |
| BIO | Biology | STEM |
| BMED | Biomedical Engineering | STEM |
| BRAE | BioResource and Agricultural Engineering | STEM |
| BUS | Business | non-STEM |
| CD | Child Development | non-STEM |
| CE | Civil Engineering | STEM |
| CHEM | Chemistry | STEM |
| CM | Construction Management | non-STEM |
| COMS | Communication Studies | non-STEM |
| CPE | Computer Engineering | STEM |
| CRP | City and Regional Planning | non-STEM |
| CRSC | Crop Science | STEM |
| CSC | Computer Science | STEM |
| DSCI | Dairy Science | STEM |
| ECON | Economics | non-STEM |
| EE | Electrical Engineering | STEM |
| EHS | Environmental Horticulture Science | STEM |
| ENGL | English | non-STEM |
| ENVE | Environmental Engineering | STEM |
| ENVM | Environmental Management and Protection | STEM |
| ERSC | Earth Sciences | STEM |
| ES | Ethnic Studies | non-STEM |
| FDSC | Food Science | STEM |
| FNR | Forestry and Natural Resources | STEM |
| FRSC | Fruit Science | STEM |
| GENE | General Engineering | STEM |
| GRC | Graphic Communication | non-STEM |
| HIST | History | non-STEM |
| IE | Industrial Engineering | STEM |
| IT | Industrial Technology | STEM |

| | | |
|------|--------------------------------------|----------|
| JOUR | Journalism | non-STEM |
| KINE | Kinesiology | STEM |
| LAES | Liberal Arts and Engineering Studies | STEM |
| LARC | Landscape Architecture | non-STEM |
| LS | Liberal Studies | non-STEM |
| MATE | Materials Engineering | STEM |
| MATH | Mathematics | STEM |
| MCRO | Microbiology | STEM |
| ME | Mechanical Engineering | STEM |
| MFGE | Manufacturing Engineering | STEM |
| MLL | Modern Languages and Literature | non-STEM |
| MU | Music | non-STEM |
| NUTR | Nutrition | STEM |
| PHIL | Philosophy | non-STEM |
| PHYS | Physics | STEM |
| POLS | Political Science | non-STEM |
| PSC | Physical Science | STEM |
| PSY | Psychology | non-STEM |
| REC | Recreation | non-STEM |
| SOCS | Social Sciences | non-STEM |
| SS | Soil Science | STEM |
| STAT | Statistics | STEM |
| TH | Theatre | non-STEM |
| WVIT | Wine and Viticulture | non-STEM |
| SE | Software Engineering | STEM |

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